

# Package ‘omicwas’

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**Type** Package

**Title** Cell-Type-Specific Association Testing in Bulk Omics Experiments

**Version** 0.8.0

**Description** In bulk epigenome/transcriptome experiments, molecular expression is measured in a tissue, which is a mixture of multiple types of cells. This package tests association of a disease/phenotype with a molecular marker for each cell type. The proportion of cell types in each sample needs to be given as input. The package is applicable to epigenome-wide association study (EWAS) and differential gene expression analysis. Takeuchi and Kato (submitted)  
``omicwas: cell-type-specific epigenome-wide and transcriptome association study".

**URL** <https://github.com/fumi-github/omicwas>

**BugReports** <https://github.com/fumi-github/omicwas/issues>

**Depends** R (>= 3.6.0)

**biocViews**

**License** GPL-3

**Encoding** UTF-8

**LazyData** true

**Imports** broom, data.table, dplyr, ff, glmnet, magrittr, MASS, matrixStats, parallel, purrr, rlang, tidyr

**RoxygenNote** 7.1.1

**Suggests** testthat, knitr, rmarkdown

**VignetteBuilder** knitr

**NeedsCompilation** no

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**Repository** CRAN

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## R topics documented:

ctassoc . . . . .	2
ctcisQTL . . . . .	5
GSE42861small . . . . .	6
GSE79262small . . . . .	7
GTExsmall . . . . .	8
rrs.fit . . . . .	9

<b>Index</b>	<b>10</b>
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ctassoc	<i>Cell-Type-Specific Association Testing</i>
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### Description

Cell-Type-Specific Association Testing

### Usage

```
ctassoc(
  X,
  W,
  Y,
  C = NULL,
  test = "full",
  regularize = FALSE,
  num.cores = 1,
  chunk.size = 1000,
  seed = 123
)
```

### Arguments

X	Matrix (or vector) of traits; samples x traits.
W	Matrix of cell type composition; samples x cell types.
Y	Matrix (or vector) of bulk omics measurements; markers x samples.
C	Matrix (or vector) of covariates; samples x covariates. X, W, Y, C should be numeric.
test	Statistical test to apply; either "full", "marginal", "nls.identity", "nls.log", "nls.logit", "propdiff.identity", "propdiff.log", "propdiff.logit" or "reducedrankridge".
regularize	Whether to apply Tikhonov (ie ridge) regularization to $\beta_{hjk}$ . The regularization parameter is chosen automatically according to an unbiased version of (Lawless & Wang, 1976). Effective for nls.* and propdiff.* tests.
num.cores	Number of CPU cores to use. Full, marginal and propdiff tests are run in serial, thus num.cores is ignored.

chunk.size	The size of job for a CPU core in one batch. If you have many cores but limited memory, and there is a memory failure, decrease num.cores and/or chunk.size.
seed	Seed for random number generation.

## Details

Let the indexes be  $h$  for cell type,  $i$  for sample,  $j$  for marker (CpG site or gene),  $k$  for each trait that has cell-type-specific effect, and  $l$  for each trait that has a uniform effect across cell types. The input data are  $X_{ik}$ ,  $C_{il}$ ,  $W_{ih}$  and  $Y_{ji}$ , where  $C_{il}$  can be omitted.  $X_{ik}$  and  $C_{il}$  are the values for two types of traits, showing effects that are cell-type-specific or not, respectively. Thus, calling  $X_{ik}$  and  $C_{il}$  as "traits" and "covariates" gives a rough idea, but is not strictly correct.  $W_{ih}$  represents the cell type composition and  $Y_{ji}$  represents the marker level, such as methylation or gene expression. For each tissue sample, the cell type proportion  $W_{ih}$  is the proportion of each cell type in the bulk tissue, which is measured or imputed beforehand. The marker level  $Y_{ji}$  in bulk tissue is measured and provided as input.

The parameters we estimate are the cell-type-specific trait effect  $\beta_{hjk}$ , the tissue-uniform trait effect  $\gamma_{jl}$ , and the basal marker level  $\alpha_{hj}$  in each cell type.

We first describe the conventional linear regression models. For marker  $j$  in sample  $i$ , the marker level specific to cell type  $h$  is

$$\alpha_{hj} + \sum_k \beta_{hjk} * X_{ik}.$$

This is a representative value rather than a mean, because we do not model a probability distribution for cell-type-specific expression. The bulk tissue marker level is the average weighted by  $W_{ih}$ ,

$$\mu_{ji} = \sum_h W_{ih} [\alpha_{hj} + \sum_k \beta_{hjk} * X_{ik}] + \sum_l \gamma_{jl} C_{il}.$$

The statistical model is

$$Y_{ji} = \mu_{ji} + \epsilon_{ji},$$

$$\epsilon_{ji} \sim N(0, \sigma_j^2).$$

The error of the marker level is normally distributed with variance  $\sigma_j^2$ , independently among samples.

The full model is the linear regression

$$Y_{ji} = \left( \sum_h \alpha_{hj} * W_{ih} \right) + \left( \sum_{hk} \beta_{hjk} * W_{ih} * X_{ik} \right) + \left( \sum_l \gamma_{jl} * C_{il} \right) + error.$$

The marginal model tests the trait association only in one cell type  $h$ , under the linear regression,

$$Y_{ji} = \left( \sum_{h'} \alpha_{h'j} * W_{ih'} \right) + \left( \sum_k \beta_{hjk} * W_{ih} * X_{ik} \right) + \left( \sum_l \gamma_{jl} * C_{il} \right) + error.$$

The nonlinear model simultaneously analyze cell type composition in linear scale and differential expression/methylation in log/logit scale. The normalizing function is the natural logarithm  $f = \log$  for gene expression, and  $f = \text{logit}$  for methylation. Conventional linear regression can be formulated by defining  $f$  as the identity function. The three models are named `nls.log`, `nls.logit` and

`nls.identity`. We denote the inverse function of  $f$  by  $g$ ;  $g = \exp$  for gene expression, and  $g = \text{logistic}$  for methylation. The mean normalized marker level of marker  $j$  in sample  $i$  becomes

$$\mu_{ji} = f\left(\sum_h W_{ih} g(\alpha_{hj} + \sum_k \beta_{hjk} * X_{ik})\right) + \sum_l \gamma_{jl} C_{il}.$$

The statistical model is

$$f(Y_{ji}) = \mu_{ji} + \epsilon_{ji},$$

$$\epsilon_{ji} \sim N(0, \sigma_j^2).$$

The error of the marker level is normally distributed with variance  $\sigma_j^2$ , independently among samples.

The ridge regression aims to cope with multicollinearity of the interacting terms  $W_{ih} * X_{ik}$ . Ridge regression is fit by minimizing the residual sum of squares (RSS) plus  $\lambda \sum_{hk} \beta_{hjk}^2$ , where  $\lambda > 0$  is the regularization parameter.

The `propdiff` tests try to cope with multicollinearity by, roughly speaking, using mean-centered  $W_{ih}$ . We obtain, instead of  $\beta_{hjk}$ , the deviation of  $\beta_{hjk}$  from the average across cell types. Accordingly, the null hypothesis changes. The original null hypothesis was  $\beta_{hjk} = 0$ . The null hypothesis when centered is  $\beta_{hjk} - (\sum_{ih'} W_{ih'} \beta_{h'jk}) / (\sum_{ih'} W_{ih'}) = 0$ . It becomes difficult to detect a signal for a major cell type, because  $\beta_{hjk}$  would be close to the average across cell types. The tests `propdiff.log` and `propdiff.logit` include an additional preprocessing step that converts  $Y_{ji}$  to  $f(Y_{ji})$ . Apart from the preprocessing, the computations are performed in linear scale. As the preprocessing distorts the linearity between the dependent variable and (the centered)  $W_{ih}$ , I actually think `propdiff.identity` is better.

## Value

A list with one element, which is named "coefficients". The element gives the estimate, statistic, p.value in tibble format. In order to transform the estimate for  $\alpha_{hj}$  to the original scale, apply `plogis` for `test = nls.logit` and `exp` for `test = nls.log`. The estimate for  $\beta_{hjk}$  by `test = nls.log` is the natural logarithm of fold-change, not the  $\log_2$ . If numerical convergence fails, NA is returned for that marker.

## References

Lawless, J. F., & Wang, P. (1976). A simulation study of ridge and other regression estimators. *Communications in Statistics - Theory and Methods*, 5(4), 307–323. <https://doi.org/10.1080/03610927608827353>

## See Also

`ctcisQTL`

## Examples

```
data(GSE42861small)
X = GSE42861small$X
W = GSE42861small$W
Y = GSE42861small$Y
```

```
C = GSE42861small$C
result = ctassoc(X, W, Y, C = C)
result$coefficients
```

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ctcisQTL

*Cell-Type-Specific QTL analysis*

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### Description

Cell-Type-Specific QTL analysis

### Usage

```
ctcisQTL(
  X,
  Xpos,
  W,
  Y,
  Ypos,
  C = NULL,
  max.pos.diff = 1e+06,
  outdir = tempdir(),
  outfile = "ctcisQTL.out.txt"
)
```

### Arguments

X	Matrix (or vector) of SNP genotypes; SNPs x samples.
Xpos	Vector of the physical position of X
W	Matrix of cell type composition; samples x cell types.
Y	Matrix (or vector) of bulk omics measurements; markers x samples.
Ypos	Vector of the physical position of Y
C	Matrix (or vector) of covariates; samples x covariates. X, Xpos, W, Y, Ypos, C should be numeric.
max.pos.diff	Maximum positional difference to compute cis-QTL. Association analysis is performed between a row of X and a row of Y, only when they are within this limit. Since the limiting is only by position, the function needs to be run separately for each chromosome.
outdir	Output directory.
outfile	Output file.

**Details**

A function for analyses of QTL, such as eQTL, mQTL, pQTL. The statistical test is almost identical to `ctassoc(test = "nls.identity", regularize = "TRUE")`. Association analysis is performed between each row of  $Y$  and each row of  $X$ . Usually, the former will be a methylation/expression marker, and the latter will be a SNP. To cope with the large number of combinations, the testing is limited to pairs whose position is within the difference specified by `max.pos.diff`; i.e., limited to cis-QTL. In detail, this function performs linear ridge regression, whereas `ctassoc(test = "nls.identity", regularize = "TRUE")` actually is nonlinear regression but with  $f = \text{identity}$  as normalizing transformation. In order to speed up computation, first, the parameters  $\alpha_{hj}$  and  $\gamma_{jl}$  are fit by ordinary linear regression assuming  $\beta_{hjk} = 0$ . Next,  $\beta_{hjk}$  are fit and tested by linear ridge regression (see documentation for `ctassoc`).

**Value**

The estimate, statistic, p.value are written to the specified file.

**See Also**

`ctassoc`

**Examples**

```
data(GSE79262small)
X   = GSE79262small$X
Xpos = GSE79262small$Xpos
W   = GSE79262small$W
Y   = GSE79262small$Y
Ypos = GSE79262small$Ypos
C   = GSE79262small$C
X   = X[seq(1, 3601, 100), ] # for brevity
Xpos = Xpos[seq(1, 3601, 100)]
ctcisQTL(X, Xpos, W, Y, Ypos, C = C)
```

---

GSE42861small

*Small Subset of GSE42861 Dataset From GEO*

---

**Description**

The dataset includes 336 rheumatoid arthritis cases and 322 controls. A subset of 500 CpG sites were randomly selected from the original EWAS dataset.

**Usage**

```
data(GSE42861small)
```

**Format**

An object of class `list` of length 4.

**Source**

[GEO](#)

**See Also**

`ctassoc`

**Examples**

```
data(GSE42861small)
X = GSE42861small$X
W = GSE42861small$W
Y = GSE42861small$Y
Y = Y[seq(1, 20), ] # for brevity
C = GSE42861small$C
result = ctassoc(X, W, Y, C = C)
result$coefficients
```

---

GSE79262small

*Small Subset of GSE79262 Dataset From GEO*

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**Description**

The dataset includes 53 samples. A subset of 737 CpG sites and 3624 SNPs within Chr1:100,000,000-110,000,000 were selected from the original EWAS dataset. DNA methylation was measured in T cells. The estimated proportion of CD4T, CD8T, NK cells are saved in `W`.

**Usage**

```
data(GSE79262small)
```

**Format**

An object of class `list` of length 6.

**Source**

[GEO](#)

**See Also**

`ctcisQTL`

**Examples**

```

data(GSE79262small)
X = GSE79262small$X
Xpos = GSE79262small$Xpos
W = GSE79262small$W
Y = GSE79262small$Y
Ypos = GSE79262small$Ypos
C = GSE79262small$C
X = X[seq(1, 3001, 100), ] # for brevity
Xpos = Xpos[seq(1, 3001, 100)]
Y = Y[seq(1, 501, 100), ]
Ypos = Ypos[seq(1, 501, 100)]
ctcisQTL(X, Xpos, W, Y, Ypos, C = C)

```

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GTEXsmall

*Small Subset of GTEx Dataset*


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**Description**

The dataset includes gene expression measured in whole blood for 389 samples. A subset of 500 genes were randomly selected from the original dataset.

**Usage**

```
data(GTEXsmall)
```

**Format**

An object of class list of length 4.

**Source**

[GTEx](#)

**See Also**

ctassoc

**Examples**

```

data(GTEXsmall)
X = GTEXsmall$X
W = GTEXsmall$W
Y = GTEXsmall$Y + 1
Y = Y[seq(1, 20), ] # for brevity
C = GTEXsmall$C
result = ctassoc(X, W, Y, C = C)
result$coefficients

```



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rrs.fit	<i>Fitting reduced-rank ridge regression with given rank and shrinkage penalty</i>
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### Description

Fitting reduced-rank ridge regression with given rank and shrinkage penalty This is a modification of rrs.fit in rrpac version 0.1-6. In order to handle extremely large  $q = \text{ncol}(Y)$ , generation of a  $q$  by  $q$  matrix is avoided.

### Usage

```
rrs.fit(Y, X, nrank = min(ncol(Y), ncol(X)), lambda = 1, coefSVD = FALSE)
```

### Arguments

Y	a matrix of response (n by q)
X	a matrix of covariate (n by p)
nrank	an integer specifying the desired rank
lambda	tuning parameter for the ridge penalty
coefSVD	logical indicating the need for SVD for the coefficient matrix in the output

### Value

S3 rrr object, a list consisting of

coef	coefficient of rrs
coef.ls	coefficient of least square
fitted	fitted value of rrs
fitted.ls	fitted value of least square
A	right singular matrix
Ad	singular value vector
nrank	rank of the fitted rrr

### References

Mukherjee, A. and Zhu, J. (2011) Reduced rank ridge regression and its kernel extensions.

Mukherjee, A., Chen, K., Wang, N. and Zhu, J. (2015) On the degrees of freedom of reduced-rank estimators in multivariate regression. *Biometrika*, 102, 457–477.

### Examples

```
Y <- matrix(rnorm(400), 100, 4)
X <- matrix(rnorm(800), 100, 8)
rfit <- rrs.fit(Y, X)
```

# Index

## \* datasets

GSE42861small, [6](#)

GSE79262small, [7](#)

GTEsmall, [8](#)

ctassoc, [2](#), [6](#)

ctcisQTL, [5](#)

GSE42861small, [6](#)

GSE79262small, [7](#)

GTEsmall, [8](#)

rrs.fit, [9](#)